

Recall the Setup:

- $y = X\theta^* + w, \quad X \in \mathbb{R}^{n \times d}, \quad n \ll d, \quad M \subseteq \bar{m} \subseteq \mathbb{R}^d$
model \bar{m}^\perp : perturbation

$$\hat{\theta} \in \operatorname{argmin}_{\theta \in \mathbb{R}^d} \{ \mathcal{L}(\theta) + \lambda R(\theta) \}$$

Here, \mathcal{L} is a smooth convex loss, R is a decomposable norm wrt (m, \bar{m}^\perp) , $\lambda \geq 2R^*(\nabla \mathcal{L}(\theta^*))$.

- Localization of error vector

$$\hat{\Delta} = \hat{\theta} - \theta^* \in \mathcal{C} = \{ \Delta \in \mathbb{R}^d : R(\Delta_{\bar{m}^\perp}) \leq 3R(\Delta_{\bar{m}}) + 4R(\theta_{\bar{m}^\perp}^*) \}$$

- RSC of \mathcal{L} over \mathcal{C}

$$\mathcal{L}(\theta^* + \Delta) \geq \mathcal{L}(\theta^*) + \nabla \mathcal{L}(\theta^*)^T \Delta + \kappa \|\Delta\|_2^2 - \tau^2(\theta^*) \quad \forall \Delta \in \mathcal{C}$$

Theorem: Under the above setup,

$$\| \hat{\theta} - \theta^* \|_2^2 \leq \underbrace{\frac{9\lambda^2}{4\kappa^2} \psi^2(\bar{m})}_{\text{tightness measure}} + \underbrace{\frac{2}{\kappa} [\tau^2(\theta^*) + 2\lambda R(\theta_{\bar{m}^\perp}^*)]}_{\text{model specification error}},$$

where

$$\psi(m) = \sup_{u \in m \setminus \{0\}} \frac{R(u)}{\|u\|_2}$$

Example: LASSO with exactly sparse model

$$\hat{\theta} \in \operatorname{argmin}_{\theta \in \mathbb{R}^d} \left\{ \underbrace{\frac{1}{2n} \|y - X\theta\|_2^2}_{\mathcal{L}(\theta)} + \lambda \underbrace{\|\theta\|_1}_{R(\theta)} \right\} \quad \text{--- (L)}$$

By taking S to be the support of θ^* with $|S| = s$,

\mathcal{C} takes the form

$$S = \operatorname{supp}(\theta^*) = \{ i : \theta_i^* \neq 0 \}$$

$$m(S) = \{ \theta : \theta_j = 0 \quad \forall j \notin S \} = \bar{m}(S)$$

$$\mathcal{C} = \{ \Delta \in \mathbb{R}^d : \|\Delta_{S^c}\|_1 \leq 3\|\Delta_S\|_2 \}$$

Want:

① \mathcal{L} is RSC over \mathcal{C} :

$$\frac{\|X\Delta\|_2^2}{n} \geq \kappa \|\Delta\|_2^2 \quad \forall \Delta \in \mathcal{E}. \quad (*)$$

② Estimate $R^*(\nabla \mathcal{L}(\theta^*))$: $R^*(0) = \|\theta\|_\infty$

$$\therefore R^*(\nabla \mathcal{L}(\theta^*)) = \frac{1}{n} \|X^T w\|_\infty$$

$$\hookrightarrow w = y - X\theta^*$$

Proposition: Suppose that

(i) X satisfies $(*)$; [RSC]

(ii) $\frac{\|X_j\|_2}{\sqrt{n}} \leq 1 \quad \forall j$, where X_j is the j^{th} col of X , [scaling]

(iii) w is sub-Gaussian with parameter $\sigma > 0$: w has zero mean and for any fixed $v \in \mathbb{R}^n$ with $\|v\|_2 = 1$,

$$\Pr[|v^T w| \geq t] \leq 2 \exp(-t^2/2\sigma^2) \quad \forall t > 0 \quad \left[\begin{array}{l} \text{Statistical} \\ \text{model of} \\ \text{noise} \end{array} \right]$$

(iv) θ^* is s -sparse. [sparsity]

Then, by taking $\lambda = 4\sigma \sqrt{\frac{\log d}{n}}$ in (L), then

$$\|\hat{\theta} - \theta^*\|_2^2 \leq 36 \frac{\sigma^2}{\kappa^2} \cdot \frac{s \log d}{n}$$

with probability $\geq 1 - 2/d$.

Proof: In the theorem, $\tau(\cdot) = 0$, $R(\theta_{\eta^\pm}^*) = 0$. It remains

to estimate $\psi(\eta)$ and λ

$$1^\circ: \quad \psi(\eta) = \sup_{u \in \eta \setminus \{0\}} \frac{\|u\|_1}{\|u\|_2} = \sqrt{s} \quad \left[\forall u \in \mathbb{R}^d: \|u\|_1 \leq \sqrt{d} \cdot \|u\|_2 \right]$$

2^o: From our calculation, we need to estimate $\frac{1}{n} \|X^T w\|_\infty$

Observe: The j^{th} entry of $X^T w$ is $X_j^T w$. Now,

$$\begin{aligned} \Pr \left[\left| \frac{X_j^T w}{n} \right| \geq t \right] &= \Pr \left[\left| \left(\frac{X_j}{\sqrt{n}} \right)^T w \right| \geq \sqrt{n} t \right] \\ &\leq 2 \exp \left(-\frac{n t^2}{2} \right) \quad \forall t > 0, \end{aligned}$$

$$\leq 2 \exp\left(-\frac{nt^2}{2\sigma^2}\right) \quad \forall t > 0, \quad j=1, \dots, d$$

By the Union bound,

$$\Pr\left[\frac{1}{n} \|X^T w\|_\infty \geq t\right] \leq 2d \exp\left(-\frac{nt^2}{2\sigma^2}\right)$$

Set $t^2 = \frac{4\sigma^2 \log d}{n}$, Then, the RHS = $\frac{2}{d}$. Hence, with probability $1 - \frac{2}{d}$, we have $\lambda = 4\sigma \sqrt{\frac{\log d}{n}} \geq \frac{2}{n} \|X^T w\|_\infty$

Establishing the RSC of $\mathcal{L}(\theta) = \frac{1}{2n} \|y - X\theta\|_2^2$ over

$$\mathcal{C} = \left\{ \Delta \in \mathbb{R}^d : \|\Delta_S\|_1 \leq 3\|\Delta_S\|_1 \right\}; \text{ i.e.,}$$

$$\frac{\|X\Delta\|_2^2}{2n} \geq k \|\Delta\|_2^2 \quad \forall \Delta \in \mathcal{C}. \quad (*)$$

Note: one would not expect (*) to hold for arbitrary X :

e.g. fix $\Delta \in \mathcal{C}$ and construct X s.t. $\Delta \in \text{null}(X)$.

→ need, e.g., a statistical model of X in order to proceed.

Theorem: Suppose that the rows of X are iid $\mathcal{N}(0, I_d)$.

Then, with probability $\geq 1 - c_1 \exp(-c_2 n)$,

where $c_1, c_2 > 0$ are constants,

$$n \begin{bmatrix} -x_j^T \\ \vdots \\ x \end{bmatrix}$$

$$(**) \quad \frac{\|X\Delta\|_2}{\sqrt{n}} \geq \frac{1}{4} \|\Delta\|_2 - 9 \sqrt{\frac{\log d}{n}} \|\Delta\|_2 \quad \forall \Delta \in \mathbb{R}^d$$

Implication: Note that for $\Delta \in \mathcal{C}$,

$$\|\Delta\|_1 = \|\Delta_S\|_1 + \|\Delta_{S^c}\|_1 \leq 4\|\Delta_S\|_1 \leq 4\sqrt{s} \|\Delta_S\|_2 \leq 4\sqrt{s} \|\Delta\|_2 \quad \Delta \in \mathcal{C}$$

Hence, (**) implies

$$\frac{\|X\Delta\|_2}{\sqrt{n}} \geq \left(\frac{1}{4} - 36 \sqrt{\frac{s \log d}{n}} \right) \|\Delta\|_2 \quad \forall \Delta \in \mathcal{G}$$

Thus, as long as

$$\frac{1}{4} - 36 \sqrt{\frac{s \log d}{n}} > 0 \iff n > 144^2 s \log d,$$

the bound is meaningful.

The proof of the theorem has 3 main steps:

Step 1: Observe that (**) is trivial if $\Delta = 0$. So we may assume $\|\Delta\|_2 = 1$ by homogeneity. For any $r > 0$, define

$$V(r) = \left\{ \Delta \in \mathbb{R}^d : \|\Delta\|_2 = 1, \|\Delta\|_1 \leq r \right\}$$

Proposition 1:

$$\mathbb{E} \left[\inf_{\Delta \in V(r)} \frac{\|X\Delta\|_2}{\sqrt{n}} \right] \geq 3 \left[\frac{1}{4} - \sqrt{\frac{\log d}{n}} r \right]$$

whenever $V(r) \neq \emptyset$.